

1 Biomarker Attenuation Following Virtual
2 Resection Validates Clinically Defined
3 Epileptogenic Networks

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20 **Abstract**

21 Patients with medically refractory epilepsy who fail antiseizure medications may
22 undergo surgical resection of the epileptogenic zone (EZ), yet postoperative out-
23 comes remain highly variable, with only 30–70% achieving long-term seizure
24 freedom. Computational virtual resection—the simulation of removing candi-
25 date brain regions from patient-specific network models—offers a principled
26 framework for evaluating proposed surgical plans prior to intervention. Because
27 clinically defined EZs are hypothesis-driven and may be incomplete or incorrect,
28 an objective test of their network impact is critically needed. Here, we assess the
29 leading eigenvector (LEV) of dynamic network models (DNMs), derived from
30 intracranial EEG (iEEG), as a biomarker for virtual resection efficacy. Using
31 iEEG recordings from 44 patients across four epilepsy centers, we constructed

32 DNMs and quantified LEV components before and after virtual resection of clinically
33 annotated EZ channels. Attenuation of EZ-associated LEV components
34 following virtual resection distinguished favorable (Engel Class 1) from unfavorable
35 (Engel Class 2–4) surgical outcomes, achieving an area under the curve
36 of 0.8, with 71% accuracy, 74% specificity, and 67% sensitivity. These results
37 establish LEV attenuation as a computational biomarker for evaluating clinically
38 defined epileptogenic network hypotheses and suggest that LEV-based virtual
39 resection may inform patient-specific surgical decision-making.

40 **Keywords:** epilepsy, virtual resection, dynamic network models, leading eigenvector,
41 interictal biomarker, epileptogenic zone, surgical outcome

42 1 Introduction

43 Epilepsy is a chronic neurological disorder characterized by recurrent, unprovoked
44 seizures that affects over 60 million people worldwide [1]. Approximately one-third of
45 individuals with epilepsy suffer from medically refractory epilepsy (MRE), defined as
46 inadequate seizure control despite treatment with two or more appropriately chosen
47 antiseizure medications [2–5]. For these patients, surgical resection of the epileptogenic
48 zone (EZ), which comprises specific brain region(s) necessary and sufficient for initiating
49 seizures and whose removal (or disconnection) results in complete abolition of
50 seizures [6]. However, surgical outcomes remain highly variable, with only 30-70% of
51 patients achieving long-term seizure freedom, largely due to challenges in accurately
52 localizing the EZ [7–9].

53 Presurgical evaluation of MRE often requires prolonged inpatient monitoring with
54 invasive intracranial EEG (iEEG), during which clinicians await spontaneous seizures
55 to visually identify early ictal (seizure) electrographic changes that delineate the
56 EZ [7, 10, 11]. This process is time-consuming, costly, and increases patient risk due to
57 extended electrode implantation. In contrast, interictal (between seizure) recordings
58 are abundant and safer to acquire, motivating efforts to identify interictal biomarkers
59 of the EZ. High-frequency oscillations (HFOs) have been proposed for this purpose,
60 but their clinical utility has been limited by inconsistent performance, challenges in
61 distinguishing pathological from physiological HFOs, and channel-wise analyses that
62 fail to capture network-level dynamics [12–14].

63 Network-based computational approaches have emerged as a promising framework
64 for understanding epilepsy as a disorder of brain connectivity rather than focal pathology
65 [15]. Patient-specific dynamic network models (DNMs) can be estimated from
66 short iEEG recordings to capture directional influences among brain regions [16]. This
67 framework enables the development of network-level biomarkers reflecting the underlying
68 dynamical properties of the epileptic brain. The source-sink index (SSI), which
69 quantifies the extent to which certain nodes are persistently inhibited by other network
70 regions during interictal periods, has shown improved predictive performance for surgical
71 outcomes compared to HFOs [17]. More recently, the leading eigenvector (LEV)
72 of the DNM state-transition matrix has been proposed as a theoretically principled

73 biomarker that is highly correlated with the SSI [18]. In linear systems theory, the
74 LEV represents the steady-state distribution of network activity, with nodes exhibit-
75 ing high LEV components corresponding to network sinks that accumulate influence
76 from other regions [19].

77 Virtual resection—the computational simulation of removing candidate brain
78 regions from patient-specific models—has emerged as a promising framework for eval-
79 uating surgical hypotheses prior to intervention [20–25]. By simulating the removal
80 of specific brain regions in patient-specific computational models, virtual resection
81 enables testing of different surgical hypotheses before actual intervention. Prior work
82 has shown that virtual resection strategies can, in principle, predict surgical outcomes
83 or identify optimal resection targets. However, existing approaches have relied on
84 small, single-center cohorts [20], required ictal or peri-ictal data [20], involved com-
85 putationally intensive neural mass models with extensive parameter tuning [26], or
86 depended on specialized neuroimaging modalities such as diffusion MRI [27]. Other
87 studies have focused primarily on optimizing resection extent rather than predicting
88 clinical outcome [21], or have been validated largely in simulation or narrowly defined
89 patient populations [24]. Collectively, these limitations have hindered translation of
90 virtual resection methods into routine clinical decision-making.

91 demonstrated that synchronizability-based virtual resection could predict surgical
92 outcomes with an AUC of 0.89, reporting that removal of synchronizing brain regions
93 was associated with improved clinical outcomes [20]. However, their approach relied on
94 a single-center cohort of 28 patients from the University of Pennsylvania and required
95 functional connectivity estimates derived from both pre-ictal and ictal recordings,
96 limiting applicability in clinical settings where waiting for seizures prolongs invasive
97 monitoring. While recent work has demonstrated the potential of computational mod-
98 eling, neural mass models require extensive parameter tuning and are computationally
99 intensive [26]. however, this approach requires specialized neuroimaging infrastructure
100 and computational resources that may limit widespread clinical adoption. Nissen et
101 al. demonstrated that optimized virtual resections on individual structural brain net-
102 works could achieve effects comparable to actual surgery while sparing on average
103 27.5% of tissue, though their study of 19 patients focused primarily on optimizing
104 resection extent rather than predicting surgical outcomes [21]. Other studies have been
105 limited by small sample sizes, Li et al. analyzed only 16 patients using autoregressive
106 connectivity networks [23], while Sun et al. validated their sEEG propagation network
107 approach on limited clinical data [25]. Kalitzin et al. validated their virtual resection
108 paradigm primarily on computational simulations rather than clinical outcomes [24].
109 Demuru et al. restricted their analysis to 11 patients with extra-temporal focal epilepsy
110 and concluded from their validation study that virtual resection methods “are not yet
111 mature to inform clinical decision-making” [22].

112 While interictal network biomarkers such as the SSI and LEV can characterize clin-
113 ically defined EZs, their utility for virtual resection—specifically, whether biomarker
114 changes induced by simulated resection predict actual surgical outcomes—remains
115 unexplored. Establishing this link is essential for transforming network-based biomark-
116 ers into actionable tools for presurgical evaluation. Here, we test the hypothesis that

117 attenuation of EZ-associated LEV components following virtual resection of clinically
118 annotated regions predicts surgical outcome. Using iEEG data from 44 MRE
119 patients across four epilepsy centers, we constructed patient-specific DNMs and quantified
120 LEV components before and after simulated resection. We show that LEV-based
121 virtual resection achieves clinically meaningful predictive accuracy, validating LEV
122 attenuation as a computational biomarker for evaluating proposed surgical plans and
123 informing patient-specific epilepsy surgery.

124 **2 Results**

125 **2.1 Virtual Resection Produces Differential LEV Changes in** 126 **Successful versus Failed Surgical Outcomes**

127 To evaluate the utility of the leading eigenvector (LEV) as a biomarker for virtual
128 resection, we examined the distribution of LEV changes across patients following simulated
129 removal of clinically annotated epileptogenic zone (EZ) channels. For each iEEG
130 channel, we computed $\Delta\text{LEV} = |\text{LEV}_{pre}| - |\text{LEV}_{post}|$, which quantifies the change in
131 LEV magnitude induced by virtual resection. Positive values indicate attenuation of
132 network influence following simulated EZ removal (Figure 1).

133 Distinct distributional patterns emerged that corresponded to surgical outcome.
134 In patients who achieved seizure freedom (Engel Class 1), EZ channels were disproportionately
135 represented in the upper tail of the ΔLEV distribution. For example, patient UPMC6
136 exhibited an EZ tail proportion of 0.784, indicating that virtual resection produced large
137 LEV reductions primarily within EZ-labeled regions (Figure 2). This pattern is consistent
138 with the hypothesis that successful resections target network elements exerting strong
139 influence on interictal dynamics, such that their removal substantially alters the network's
140 steady-state behavior.

141 In contrast, patients with unfavorable surgical outcomes showed markedly lower EZ
142 representation in the distribution tail. Patient NIH9 (Engel Class 3) exhibited an EZ
143 tail proportion of 0.175, while patient LA15 (Engel Class 4) showed a value of 0.078.
144 In these cases, virtual resection of clinically identified EZ channels produced minimal
145 changes in LEV magnitude, suggesting that the annotated EZ either incompletely
146 captured the true epileptogenic network or omitted additional regions critical to seizure
147 generation.

148 Kernel density estimates overlaid on the ΔLEV histograms (Figure 2) revealed
149 that the overall shape of the distributions was broadly similar across patients. The primary
150 distinguishing feature was the composition of extreme values rather than global
151 distributional differences. This observation supports the use of tail-based metrics as
152 a standardized measure across patients and recording conditions. Complete ΔLEV
153 distributions for all 44 patients are provided in Supplementary Figures S1–S4.

154 **2.2 EZ Tail Proportion Differentiates Surgical Outcome** 155 **Groups**

156 We next asked whether the proportion of EZ channels in the upper tail of the ΔLEV
157 distribution could serve as a patient-level predictor of surgical outcome. For each

158 patient, we computed the fraction of channels within the top 5% of ΔLEV values that
159 originated from clinically annotated EZ regions, hereafter referred to as the “EZ tail
160 proportion”.

161 Patients achieving Engel Class 1 outcomes ($n = 21$) exhibited substantially higher
162 EZ tail proportions than those with Engel Class 2–4 outcomes ($n = 23$) (Figure
163 3). Logistic regression confirmed a significant association between EZ tail proportion
164 and surgical outcome (likelihood ratio test, $p = 4.49 \times 10^{-5}$). Leave-one-out cross-
165 validation demonstrated robust predictive performance with an accuracy of 71% and
166 an area under the receiver operating characteristic curve (AUC) of 0.8. These results
167 reflect the central premise of our approach: when the clinically defined EZ accurately
168 captures the epileptogenic network, virtual resection induces large ΔLEV reductions
169 concentrated within EZ channels. These findings support the EZ tail proportion as a
170 candidate biomarker linking virtual resection to clinical outcome.

171 **2.3 LEV-Based Virtual Resection Enables Clinically** 172 **Meaningful Outcome Prediction**

173 To formally assess predictive performance, we trained a logistic regression classifier
174 using EZ tail proportion as the sole predictor for binary surgical outcome (Engel Class
175 1 versus Engel Class 2–4). Model performance was evaluated using leave-one-out cross-
176 validation across all 44 patients (21 Engel Class 1; 23 Engel Class 2–4), with each
177 patient iteratively held out for testing while the model was trained on the remaining
178 cohort.

179 The classifier achieved robust and consistent performance across folds. On the held-
180 out test cases, the model yielded an area under the receiver operating characteristic
181 curve (AUC) of 0.8, with 71% accuracy, 74% specificity, and 67% sensitivity (Figure 4).
182 The receiver operating characteristic curve demonstrated stable discrimination across
183 a range of decision thresholds. Analysis of model coefficients confirmed the statistical
184 significance of EZ tail proportion as a predictor ($p = 0.0014$; Table 2). The positive
185 standardized coefficient indicates that higher EZ tail proportion values are associated
186 with increased likelihood of seizure freedom, with each unit increase corresponding to
187 approximately a 1.5-fold increase in the odds of achieving an Engel Class 1 outcome.

188 The resulting confusion matrix revealed six false-positive predictions (predicted
189 Engel Class 1, observed Engel Class 2–4) and seven false-negative predictions (pre-
190 dicted Engel Class 2–4, observed Engel Class 1), highlighting clinically interpretable
191 error modes and underscoring the model’s role as a decision-support tool rather than
192 a deterministic predictor.

193 **2.4 LEV Performance is Invariant to Demographic and** 194 **Clinical Characteristics**

195 To assess the clinical generalizability of the LEV-based EZ localization metric, we
196 evaluated whether the EZ tail proportion was confounded by demographic or clinical
197 factors unrelated to surgical outcome. Following previously established methodol-
198 ogy for assessing biomarker invariance [16], we computed Cohen’s d effect sizes with

199 95% bootstrap confidence intervals and performed non-parametric group comparisons
200 across key clinical covariates (Figure 5).

201 No significant difference in EZ tail proportion was observed between male ($n = 23$;
202 mean \pm SD: 0.318 ± 0.206) and female ($n = 20$; mean \pm SD: 0.330 ± 0.224) patients
203 (Mann-Whitney U test, $p = 0.894$). The effect size was negligible (Cohen’s $d = -0.057$;
204 95% CI: $[-0.674, 0.561]$), indicating that LEV-based EZ localization is invariant to
205 patient sex (Figure 5A,F).

206 Patients with MRI-positive lesional epilepsy ($n = 12$; mean \pm SD: 0.316 ± 0.139)
207 and MRI-negative nonlesional epilepsy ($n = 31$; mean \pm SD: 0.327 ± 0.238) exhibited
208 comparable EZ tail proportion distributions (Mann-Whitney U test, $p = 0.797$). The
209 negligible effect size (Cohen’s $d = -0.051$; 95% CI: $[-0.738, 0.635]$) confirms that the
210 LEV metric is equally effective for both lesional and nonlesional cases (Figure 5C,F).
211 This finding is particularly clinically relevant, as nonlesional epilepsy cases represent
212 a challenging population where traditional imaging-based EZ localization often fails.

213 Comparison between temporal lobe epilepsy ($n = 5$; mean \pm SD: 0.409 ± 0.189) and
214 extra-temporal epilepsy ($n = 26$; mean \pm SD: 0.320 ± 0.200) revealed a small effect size
215 (Cohen’s $d = 0.431$; 95% CI: $[-0.574, 1.436]$) that did not reach statistical significance
216 (Mann-Whitney U test, $p = 0.658$). While the wide confidence interval reflects the
217 limited temporal lobe sample size, the non-significant difference suggests the LEV
218 metric performs consistently across different epileptogenic network topologies (Figure
219 5D,F).

220 Across four independent clinical centers (CC, JHU, NIH, UPMC), no significant
221 overall difference in EZ tail proportion was detected (Kruskal-Wallis test, $H = 1.447$,
222 $p = 0.695$). However, the maximum pairwise effect size between centers (CC vs.
223 UPMC: Cohen’s $d = 0.546$) fell within the medium effect range, suggesting modest
224 inter-center variability (Figure 5B,E). Importantly, despite this variability, the sep-
225 aration between good (Engel I) and poor (Engel II+) surgical outcomes remained
226 consistent across all centers (Figure 5E), with Engel I patients consistently demon-
227 strating higher EZ tail proportions than Engel II+ patients at each site. This
228 pattern indicates that while absolute EZ tail proportion values may vary slightly by
229 center—potentially reflecting differences in electrode placement protocols, data acqui-
230 sition systems, or patient selection criteria—the relative discriminative capacity of the
231 LEV metric is preserved.

232 As a positive control, we confirmed that surgical outcome itself produced a large
233 effect on EZ tail proportion, as expected if the metric captures clinically meaningful
234 information. Engel I patients ($n = 21$; mean \pm SD: 0.449 ± 0.194) exhibited significantly
235 higher EZ tail proportions than Engel II+ patients ($n = 23$; mean \pm SD: 0.206 ± 0.155 ;
236 Mann-Whitney U test, $p < 0.001$). The large effect size (Cohen’s $d = 1.36$; 95%
237 CI: $[0.684, 2.036]$) validates that the LEV-based metric captures variance specifically
238 associated with surgical outcome rather than demographic or clinical confounders
239 (Figure 5F).

240 Taken together, these results demonstrate that LEV-based EZ localization is robust
241 to patient sex and MRI lesional status (negligible effects, $|d| < 0.2$), with modest
242 sensitivity to epilepsy localization type ($|d| = 0.43$) and medical center ($|d| = 0.55$).
243 Critically, the large effect size for surgical outcome ($d = 1.36$) relative to demographic

244 factors confirms that the observed differences in EZ tail proportion reflect true predic-
245 tive signal rather than systematic bias from clinical covariates. These findings support
246 the generalizability of the LEV biomarker across heterogeneous patient populations
247 and clinical settings.

248 **3 Discussion**

249 In this study, we evaluated the leading eigenvector (LEV) of patient-specific dynamic
250 network models as a biomarker for virtual resection, that is the computational test-
251 ing of proposed surgical plans, in drug-resistant epilepsy. We show that attenuation of
252 EZ-associated LEV components following simulated resection of clinically annotated
253 regions reliably distinguishes favorable from unfavorable surgical outcomes in a mul-
254 ticenter cohort. These findings establish LEV-based virtual resection as a principled,
255 network-level approach for evaluating presurgical hypotheses using interictal data.

256 **3.1 Leading Eigenvector as a Perturbation-Responsive 257 Biomarker**

258 The utility of the LEV as a biomarker is grounded in linear systems theory, where the
259 leading eigenvector reflects the dominant steady-state mode of a dynamical network
260 and captures how activity is distributed across nodes under persistent interactions. In
261 the context of brain networks, the LEV encodes a global hierarchy of influence: nodes
262 with large positive LEV components occupy privileged positions in the network, exert-
263 ing sustained control over network dynamics, while nodes with smaller components
264 are primarily shaped by incoming influence. As such, the LEV provides a compact,
265 model-derived summary of long-timescale network organization that is not apparent
266 from pairwise connectivity measures alone.

267 Prior work established that the LEV is closely related to the source-sink index
268 (SSI), an interictal biomarker reflecting the hypothesis that epileptogenic regions
269 are persistently suppressed by surrounding networks between seizures [17]. In this
270 framework, epileptogenic zones occupy sink-like positions in the interictal state, char-
271 acterized by strong incoming influence and limited outgoing control. The LEV captures
272 this organization naturally: sink-like nodes exhibit elevated LEV magnitude because
273 they are tightly embedded in, and stabilized by, the dominant network mode. Import-
274 antly, this interpretation does not rely on transient seizure activity, but instead reflects
275 stable network structure estimated from interictal recordings.

276 Our results extend this conceptual framework in a critical direction by demonstrat-
277 ing that the LEV is not merely descriptive, but intervention-sensitive. When clinically
278 annotated EZ regions are virtually resected, implemented as targeted removal of nodes
279 or edges from the patient-specific dynamical model, the resulting perturbation propa-
280 gates through the network and alters its dominant eigenmode. We find that successful
281 surgical plans are characterized by a marked attenuation of EZ-associated LEV com-
282 ponents following virtual resection, whereas unsuccessful plans leave the dominant
283 mode largely intact. This distinction suggests that favorable outcomes occur when
284 surgery disrupts the core network structure sustaining pathological dynamics, rather
285 than simply removing regions that are anatomically or visually suspicious.

286 From a systems perspective, this behavior is expected: the leading eigenvector
287 governs the asymptotic behavior of linear dynamical systems, and perturbations that
288 significantly reduce its magnitude or redistribute its mass correspond to fundamental
289 changes in network stability and control. Thus, LEV attenuation provides a principled,
290 mechanistic criterion for evaluating whether a proposed resection meaningfully alters
291 the pathological network state. Unlike heuristics based on lesion size, spatial over-
292 lap, or local connectivity strength, the LEV directly tests whether the hypothesized
293 epileptogenic network has been functionally dismantled.

294 **3.2 Virtual Resection Approaches**

295 Our approach differs from prior virtual resection studies in several important
296 respects. A range of computational frameworks have previously been proposed to
297 evaluate surgical hypotheses, including synchronizability-based metrics derived from
298 functional connectivity [20], neural mass models that assess proximity to ictal bifur-
299 cations [26, 27], and approaches that rely on perturbations of structural connectivity
300 inferred from diffusion imaging [21]. These methods have provided valuable insight
301 into seizure generation and propagation, but often require ictal or peri-ictal data, com-
302 putationally intensive simulations, or multimodal imaging that may not be available
303 or reliable in all clinical settings.

304 In contrast, our framework operates entirely on interictal iEEG and leverages
305 steady-state properties of patient-specific dynamical network models. Rather than esti-
306 mating proximity to seizure onset or optimizing network synchronizability, LEV-based
307 virtual resection evaluates whether a proposed intervention disrupts the dominant
308 network mode that stabilizes pathological dynamics between seizures. This distinc-
309 tion is critical: whereas prior approaches often assess whether a network is rendered
310 less seizure-prone in principle, the LEV explicitly tests whether the hypothesized
311 epileptogenic network has been functionally dismantled by the proposed resection.

312 A further advantage of the LEV framework is its simplicity and interpretability.
313 The proportion of clinically annotated EZ channels in the upper tail of the Δ LEV
314 distribution yields a single scalar metric per patient that can be directly compared
315 across surgical hypotheses. This avoids the need for extensive parameter tuning or
316 model calibration and facilitates integration into clinical workflows. Despite this re-
317 lative simplicity, predictive performance was comparable to that reported by more
318 complex approaches. For example, Kini et al. reported high predictive accuracy using
319 synchronizability-based virtual resection in a smaller single-center cohort [20], while
320 Goodfellow et al. achieved similar performance using neural mass models and bifurca-
321 tion analysis [26]. Importantly, our results were validated across a multicenter cohort
322 spanning four epilepsy centers, enhancing generalizability and mitigating center-
323 specific biases that can influence model performance [22]. Together, these findings
324 suggest that steady-state network perturbations captured by the leading eigenvector
325 encode outcome-relevant information that is robust across recording sites and clini-
326 cal practices, supporting LEV-based virtual resection as a scalable and data-efficient
327 alternative for presurgical evaluation.

3.3 Relevance to Basic and Translational Neuroscience

Beyond its clinical implications, this work has direct relevance to both basic and translational neuroscience by linking dynamical systems theory to experimentally observable brain network organization. At a fundamental level, the leading eigenvector provides a principled representation of how large-scale neural networks self-organize under persistent interactions, capturing a dominant mode that reflects long-timescale stability, hierarchy, and control. Our results demonstrate that this dominant mode is not merely a mathematical abstraction, but corresponds to biologically meaningful structure that constrains pathological brain dynamics in epilepsy.

From a basic neuroscience perspective, the findings support the view that epileptogenic regions are embedded within, and stabilized by, distributed networks during the interictal state, rather than acting as autonomous seizure generators. The sensitivity of the LEV to targeted perturbations provides evidence that seizure propensity is governed by global network organization, offering a mechanistic bridge between cellular-level excitability and systems-level seizure emergence. More broadly, this work illustrates how steady-state network modes can serve as interpretable descriptors of brain function in nonstationary, pathological regimes.

Translationally, the LEV framework exemplifies how mechanistic network models can be leveraged to generate testable, intervention-relevant predictions from human neural recordings. By enabling virtual resection using interictal data, this approach transforms descriptive biomarkers into actionable computational assays, narrowing the gap between theoretical neuroscience and clinical decision-making. The same modeling principles may generalize to other network disorders—including movement disorders, chronic pain, and neuropsychiatric disease—where targeted perturbations aim to reshape maladaptive network states. Together, these results highlight the broader utility of eigenmode-based representations as a unifying language across basic, translational, and clinical neuroscience, demonstrating how abstract systems-level constructs can yield concrete insight into brain dysfunction and its remediation.

3.4 Clinical Impact

More broadly, these findings position the LEV as a bridge between network neuroscience and surgical decision-making. By enabling virtual resection, the *in silico* testing of competing surgical hypotheses, the LEV framework allows clinicians to ask not only where seizures arise, but whether a planned intervention is sufficient to neutralize the underlying pathological network. This reframes presurgical evaluation from a problem of localization alone to one of mechanism-based validation, grounded in patient-specific network dynamics.

A key clinical advantage of the LEV-based approach is its reliance on interictal data, which are routinely collected and substantially more abundant than seizures in intracranial EEG monitoring. This reduces dependence on capturing representative seizures, shortens monitoring duration, and mitigates risks associated with prolonged invasive recordings. In practice, LEV-based virtual resection could be incorporated

369 alongside standard presurgical workups to evaluate candidate resections early, provid-
370 ing objective feedback on whether a proposed plan is likely to meaningfully disrupt
371 the epileptogenic network before irreversible intervention.

372 Importantly, our results suggest that surgical failure may often arise not from incor-
373 rect localization per se, but from incomplete disruption of the pathological network. In
374 unsuccessful cases, clinically annotated EZ regions may be removed, yet the dominant
375 network mode, captured by the LEV, remains largely intact. From a clinical stand-
376 point, this offers a quantitative explanation for why anatomically plausible resections
377 sometimes fail and highlights the potential value of expanding, refining, or re-targeting
378 surgical plans to address network-level pathology rather than isolated regions.

379 The LEV framework may also support comparative evaluation of alternative inter-
380 ventions, including extended resections, staged procedures, or focal neuromodulation
381 strategies. Because the metric explicitly quantifies how proposed perturbations alter
382 the dominant network mode, it provides a common language for assessing the expected
383 impact of diverse treatment options on pathological dynamics. This is particularly rel-
384 evant in complex cases involving multifocal epilepsy, eloquent cortex, or deep network
385 hubs, where surgical decisions must balance efficacy against functional risk.

386 Finally, by grounding outcome prediction in interpretable systems-level principles
387 rather than black-box classification, LEV-based virtual resection aligns with the needs
388 of clinical decision-making. The directionality and magnitude of LEV attenuation
389 offer an intuitive measure of whether a hypothesized epileptogenic network has been
390 functionally dismantled, facilitating communication between clinicians, engineers, and
391 patients. Together, these features suggest that LEV-based virtual resection has the
392 potential to enhance surgical planning, reduce failed interventions, and move epilepsy
393 surgery toward a more principled, network-informed paradigm.

394 **3.5 Methodological Considerations and Limitations**

395 Several methodological considerations inform the interpretation of these results. First,
396 the LEV-based virtual resection framework evaluates the impact of perturbing candi-
397 date resection plans within the set of brain regions sampled during intracranial
398 monitoring. Within this sampled network, the approach enables systematic comparison
399 of alternative surgical hypotheses and can, in principle, identify resection plans that
400 more effectively disrupt the dominant pathological network mode than the clinically
401 proposed plan. However, the framework is inherently limited by electrode coverage:
402 if critical components of the epileptogenic network are not sampled by the implanted
403 electrodes, virtual resection cannot evaluate or compensate for their omission. Thus,
404 LEV-based virtual resection is well suited for refining and testing surgical strategies
405 within observed networks, but it cannot recover epileptogenic regions that lie outside
406 the monitored cortical and subcortical territory.

407 Second, the network models used here are estimated from interictal recordings and
408 capture steady-state properties of pathological dynamics. While this design choice
409 enables data-efficient modeling and avoids reliance on seizure capture, it necessarily
410 abstracts away transient ictal mechanisms and nonlinear state transitions that may
411 also contribute to surgical outcome. Thus, LEV-based virtual resection should be
412 viewed as complementary to seizure-based analyses rather than a replacement.

413 Third, outcome classification was based on seizure status at one year postop-
414 eratively. Although this time point is standard in surgical outcome studies, longer
415 follow-up may reveal delayed seizure recurrence, particularly in cases where residual
416 network pathology persists. In addition, the cohort consisted exclusively of patients
417 who proceeded to surgery, introducing potential selection bias that may limit gener-
418 alizability to earlier stages of presurgical evaluation. Prospective validation applied
419 prior to surgical decision-making will be critical for establishing clinical utility. The
420 observed false-positive and false-negative predictions further reflect these methodolog-
421 ical constraints. False positives may arise when the epileptogenic network is correctly
422 identified but cannot be fully resected due to involvement of eloquent cortex, while
423 false negatives may occur when effective resections extend beyond narrowly defined EZ
424 boundaries or when model estimation is limited by sparse sampling. These error modes
425 highlight that LEV-based virtual resection is best interpreted as a decision-support
426 tool that provides quantitative guidance rather than deterministic predictions.

427 Future work will address these limitations by prospectively evaluating multiple
428 alternative resection hypotheses *in silico*, integrating multimodal information such as
429 structural connectivity and imaging, and extending the framework to neuromodulation
430 contexts, including responsive neurostimulation. In such settings, identifying network
431 elements whose perturbation maximally attenuates the dominant pathological mode
432 may inform stimulation targeting and adaptive control strategies.

433 4 Methods

434 4.1 Experimental Model and Study Participant Details

435 4.1.1 Human Participants

436 This retrospective study analyzed 44 adults with medically refractory epilepsy
437 who underwent presurgical evaluation with intracranial stereo-electroencephalography
438 (sEEG) monitoring followed by surgical treatment. Participants ranged in age from 16
439 to 65 years (mean 31.8 ± 12.3 years) and included 24 males and 20 females. All par-
440 ticipants had at least one year of postoperative follow-up to assess clinical outcomes.
441 Participants were treated across four epilepsy centers: Cleveland Clinic (CC; $n = 27$),
442 Johns Hopkins Hospital (JHH; $n = 4$), National Institutes of Health (NIH; $n = 8$),
443 and University of Pittsburgh Medical Center (UPMC; $n = 5$). Post-iEEG interven-
444 tions included resective surgery ($n = 29$) or laser ablation ($n = 15$). Of the 44 patients,
445 21 achieved successful surgical outcomes (Engel Class 1) and 23 had failed outcomes
446 (Engel Class 2–4) at one year postoperatively. Each participating institution obtained
447 approval from their respective Institutional Review Board for retrospective analysis
448 of clinical data. All patients provided informed consent for use of their clinical data
449 in research as part of their clinical care agreements. All methods were performed in
450 accordance with the relevant guidelines and regulations. Detailed demographic and
451 clinical characteristics are summarized in Table 1.

452 **4.1.2 Sex and Gender Considerations**

453 The study cohort included 24 male and 20 female participants. Sex was determined
454 based on clinical records. Given the retrospective nature of the study and the focus on
455 computational biomarker validation, we did not perform sex-stratified analyses. The
456 influence of sex on LEV-based virtual resection metrics represents an important area
457 for future investigation in larger cohorts.

458 **4.2 Stereo-Electroencephalography Recording**

459 Intracranial stereo-EEG (sEEG) data were recorded using Nihon Kohden or Natus
460 (Natus Medical Inc.) EEG monitoring and diagnostic systems. The majority of record-
461 ings were acquired at sampling frequencies of 1 or 2 kHz, with a small subset recorded
462 at 500/512 Hz. The clinical team at each center determined electrode placement based
463 on presurgical hypotheses generated from noninvasive evaluation including scalp EEG,
464 neuroimaging, and neuropsychological testing. For each patient, a single randomly
465 selected interictal epoch was extracted, defined as occurring at least one hour away
466 from any recorded seizure event. The mean duration of analyzed epochs was $5.3 \pm$
467 4.2 minutes. No specific selection criteria were applied beyond the interictal timing
468 requirement, ensuring that the analysis reflects typical interictal network dynamics
469 rather than periictal states.

470 **4.3 Clinical Annotation of the Epileptogenic Zone**

471 The clinical team at each participating epilepsy center developed an EZ hypothesis for
472 each patient as part of standard presurgical evaluation. The clinically annotated EZ
473 was defined as the anatomical region(s) identified for targeted treatment, including
474 resection, ablation, or stimulation. This annotation was performed prior to surgery
475 based on comprehensive multimodal assessment. The EZ hypothesis included sEEG
476 channels displaying early electrophysiological changes at seizure onset, typically char-
477 acterized by low-voltage fast activity, as well as channels involved in early seizure
478 propagation. Because surgical treatment is guided by the EZ hypothesis, substan-
479 tial overlap typically exists between clinically annotated EZ channels and the actual
480 resected or ablated tissue.

481 **4.4 Classification of Surgical Outcomes**

482 Surgical outcomes were classified by each center’s epileptologist using the Engel Sur-
483 gical Outcome Scale at a minimum of 12 months postoperatively. Successful surgical
484 outcomes were defined as Engel Class 1 (free of disabling seizures), while failed out-
485 comes were defined as Engel Class 2–4 (any ongoing disabling seizures). Of the 44
486 patients in the cohort, 21 (47.7%) achieved successful outcomes and 23 (52.3%) had
487 failed outcomes.

488 **4.5 Signal Preprocessing**

489 Intracranial sEEG recordings were preprocessed using standardized procedures imple-
490 mented in MATLAB R2020b (MathWorks, Natick, MA). Each signal was bandpass

491 filtered between 0.5 Hz and 300 Hz using a fourth-order Butterworth filter to
492 retain physiologically relevant frequency content while removing DC offset and high-
493 frequency noise. Power line interference was attenuated by applying notch filters at 60
494 Hz and its harmonics (120 Hz, 180 Hz, 240 Hz) with a stopband width of 2 Hz. A com-
495 mon average reference was then computed and subtracted from all channels to reduce
496 common-mode noise and reference-related artifacts. Electrode locations were deter-
497 mined by co-registering post-implantation computed tomography (CT) scans with
498 preoperative brain MRI using BioImage Suite software. The clinical team visually ver-
499 ified electrode placement, and channels not recording from gray matter or deemed
500 excessively noisy by clinical review were excluded from analysis. This resulted in 95
501 ± 32 (mean \pm SD) sEEG channels per patient included in the final analysis. The pre-
502 processed sEEG recordings were segmented into non-overlapping 500 ms windows for
503 subsequent dynamic network model estimation. This window duration was selected to
504 balance temporal resolution with statistical stability of model parameter estimates.

505 4.6 Dynamic Network Model Estimation

506 Dynamic network models (DNMs) were used to characterize directional influences
507 among iEEG channels during interictal periods. The interictal DNM was represented
508 as a linear time-varying (LTV) model comprising a sequence of linear time-invariant
509 (LTI) models estimated from successive 500 ms data windows. For each window, we
510 estimated an LTI model of the form:

$$x(t+1) = Ax(t) \tag{1}$$

511 where $x(t) \in \mathbb{R}^N$ denotes the vector of iEEG channel amplitudes at discrete time
512 t , and $A \in \mathbb{R}^{N \times N}$ is the state-transition matrix describing the directed influences
513 between channels. Each element A_{ij} quantifies the influence of channel j on channel
514 i at the subsequent time step. This formulation can be viewed as a first-order multi-
515 variate autoregressive (MVAR) model, a widely used approach for analyzing effective
516 connectivity in brain networks.

517 The state-transition matrix A was estimated using ordinary least squares regres-
518 sion. Specifically, for each 500 ms window containing T time samples, we formed
519 the matrices $X_0 = [x(1), x(2), \dots, x(T-1)]$ and $X_1 = [x(2), x(3), \dots, x(T)]$, then
520 computed $A = X_1 X_0^T (X_0 X_0^T)^{-1}$. Model quality was assessed using the coefficient of
521 determination (R^2) between predicted and observed signals. Only A matrices achieving
522 $R^2 > 0.6$ were retained for subsequent analysis, ensuring that the estimated network
523 models adequately captured the temporal dynamics of the recorded signals. Model
524 estimation and all subsequent analyses were implemented in Python 3.11 (Python
525 Software Foundation).

526 4.7 Leading Eigenvector Computation

527 For each retained state-transition matrix A , we computed the leading eigenvector
528 (LEV), the eigenvector associated with the eigenvalue of largest magnitude. In linear
529 systems theory, the LEV characterizes the dominant mode of system dynamics and

530 represents the direction along which the system state grows most rapidly (for eigen-
531 values > 1) or decays most slowly (for eigenvalues < 1). The LEV was computed using
532 standard eigenvalue decomposition algorithms (`numpy.linalg.eig` in Python). For each
533 channel i , the corresponding LEV component v_i quantifies that channel’s participa-
534 tion in the dominant network mode. Channels with high LEV magnitude correspond
535 to network “sinks” that accumulate influence from other regions, consistent with the
536 theoretical interpretation of the source-sink framework.

537 4.8 Virtual Resection Procedure

538 Virtual resection was performed by simulating the removal of clinically annotated EZ
539 channels from the network model. Specifically, for each retained A matrix, we created
540 a post-resection matrix A_{post} by setting the columns corresponding to EZ channels to
541 zero:

$$A_{post}(:, j) = 0 \quad \text{for all } j \in \text{EZ} \quad (2)$$

542 This operation eliminates the outgoing influence of EZ channels on the rest of the
543 network while preserving their incoming connections, modeling the effect of surgical
544 removal on network dynamics. The procedure was applied independently to each A
545 matrix, yielding paired pre-resection and post-resection matrices for each 500 ms win-
546 dow. The LEV was then recomputed for each post-resection matrix A_{post} , yielding
547 LEV_{post} . The change in LEV components was quantified as:

$$\Delta\text{LEV} = |\text{LEV}_{pre}| - |\text{LEV}_{post}| \quad (3)$$

548 where the absolute value ensures that the metric captures magnitude changes
549 regardless of eigenvector sign ambiguity.

550 4.9 Feature Extraction

551 A single patient-level feature was derived from the ΔLEV data to serve as a predictor
552 of surgical outcome. For each patient, all ΔLEV values across time windows and
553 channels were pooled into a single distribution. The high-magnitude tail was defined as
554 the top 5% of values in this distribution, an exploratory threshold selected to capture
555 channels exhibiting the largest changes following virtual resection. We then calculated
556 the proportion of elements in this high-magnitude tail that originated from clinically
557 annotated EZ channels:

$$\text{EZ tail proportion} = \frac{\text{Number of tail elements from EZ channels}}{\text{Total number of tail elements}} \quad (4)$$

558 This scalar feature quantifies the extent to which virtual resection specifically
559 affects EZ channels relative to non-EZ channels. Higher values indicate that EZ chan-
560 nels disproportionately contribute to large ΔLEV changes, suggesting that the clinical
561 EZ hypothesis captures the network regions most influential in interictal dynamics.

562 4.10 Quantification and Statistical Analysis

563 4.10.1 Statistical Comparisons

564 Differences in EZ tail proportion between surgical outcome groups (Engel Class 1
565 versus Engel Class 2–4) were assessed using Mann-Whitney U tests, a non-parametric
566 alternative appropriate for non-normally distributed data and modest sample sizes.
567 Significance was defined as $p < 0.05$.

568 4.10.2 Effect Size Estimation and Invariance Assessment

569 To evaluate the clinical generalizability of the LEV-based biomarker, we assessed
570 its invariance to demographic and clinical covariates following previously established
571 methodology [16]. Effect sizes were quantified using Cohen’s d , with 95% confidence
572 intervals estimated via bootstrap resampling (10,000 iterations). For comparisons
573 across more than two groups (e.g., medical centers), the Kruskal-Wallis H test was
574 employed. Covariates evaluated included patient sex, MRI lesional status (lesional vs.
575 nonlesional), epilepsy localization (temporal vs. extra-temporal), and medical center.
576 As a positive control, we verified that surgical outcome (Engel Class 1 vs. 2+) pro-
577 duced a large effect size, confirming that the metric captures clinically meaningful
578 variance rather than demographic confounders.

579 4.10.3 Predictive Modeling

580 Logistic regression was used for binary classification of surgical outcomes. The model
581 was trained using the EZ tail proportion as the sole predictor variable, with standard-
582 ization (z-scoring) applied prior to model fitting. Model performance was evaluated
583 using multiple metrics: area under the receiver operating characteristic curve (AUC),
584 accuracy, sensitivity (true positive rate), specificity (true negative rate), and F1-score.
585 Ten-fold cross-validation was performed on the training set to assess model stabil-
586 ity and select hyperparameters. For aggregate performance assessment, leave-one-out
587 cross-validation (LOOCV) was performed across all 44 patients (21 Engel Class 1, 23
588 Engel Class 2-4). In LOOCV, each patient was iteratively held out as a single-sample
589 test set while the model was trained on the remaining 43 patients. Predictions were
590 aggregated across all folds to compute overall performance metrics. Odds ratios and
591 95% confidence intervals were computed from the exponentiated coefficients.

592 Declarations

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- 595 • **Conflict of interest:** S.V.S. is President and Co-Founder of NeuroLogic Solutions.
596 G.K. is Director of Software Development of NeuroLogic Solutions. The authors
597 declare no other competing interests.
- 598 • **Ethics approval and consent to participate:** All data were acquired with the
599 approval of the local Institutional Review Board (IRB) at each clinical institution:
600 Cleveland Clinic by the Cleveland Clinic Institutional Review Board, Johns Hopkins

Table 1 Summary of Patient Demographics and Clinical Characteristics

Characteristic	CC	JHH	NIH	UPMC	Total
Number of Patients	27	4	8	5	44
Sex (Male/Female)	13/14	2/2	6/2	3/2	24/20
Age Range (years)	16–65	24–62	16–46	23–46	16–65
Surgical Outcome					
Engel Class 1 (Success)	11	3	4	3	21
Engel Class 2–4 (Failure)	16	1	4	2	23
MRI Findings					
Normal	24	0	4	4	32
Abnormal	3	4	4	1	12

CC = Cleveland Clinic; JHH = Johns Hopkins Hospital; NIH = National Institutes of Health; UPMC = University of Pittsburgh Medical Center. Surgical outcomes were assessed at ≥ 12 months postoperatively using the Engel Surgical Outcome Scale. Success was defined as Engel Class 1 (free of disabling seizures).

Table 2 Logistic Regression Model Coefficients

Feature	Coefficient	Std. Error	z-statistic	p-value
EZ Tail Proportion	1.59	0.50	3.20	0.0014
Intercept	-0.044	0.37	-0.12	0.91

Logistic regression coefficients for predicting Engel Class 1 outcome. Final model trained on data from all 44 patients. Positive coefficients indicate that higher EZ tail proportion values are associated with increased probability of surgical success.

601 Hospital by the Johns Hopkins IRB (Protocol Number: NA_00009990; Principal
602 Investigator: Joon Kang), National Institutes of Health by the National Institutes
603 of Health IRB, and University of Pittsburgh Medical Center by the University of
604 Pittsburgh Medical Center IRB. Informed consent was obtained at each clinical
605 center. The acquisition of data for research purposes was completed with no impact
606 on the clinical objectives of the patient stay. Digitized data were stored in an IRB-
607 approved database compliant with Health Insurance Portability and Accountability
608 Act (HIPAA) regulations.

- 609 • **Consent for publication:** Not applicable. The data presented in this manuscript
610 are de-identified and aggregated; no personally identifiable information, images, or
611 videos are included.
- 612 • **Data availability:** Deidentified iEEG data reported in this study cannot be
613 deposited in a public repository due to patient privacy restrictions. Data may be
614 available upon reasonable request to S.V.S., subject to institutional data sharing
615 agreements.
- 616 • **Code availability:** All original code for dynamic network model estimation, virtual
617 resection analysis, and statistical modeling has been deposited at [GitHub repos-
618 itory to be provided upon publication] and is publicly available as of the date of
619 publication.

620 • **Author contribution:** Conceptualization, S.V.S.; Methodology, D.W. and S.V.S.;
621 Software, D.W.; Formal Analysis, D.W.; Investigation, D.W. and S.V.S.; Data Cura-
622 tion, K.M.G.; Visualization, D.W. and S.V.S.; Writing – Original Draft, D.W.,
623 E.A.P., G.K., and S.V.S.; Writing – Review & Editing, D.W., E.A.P., G.K., V.B.,
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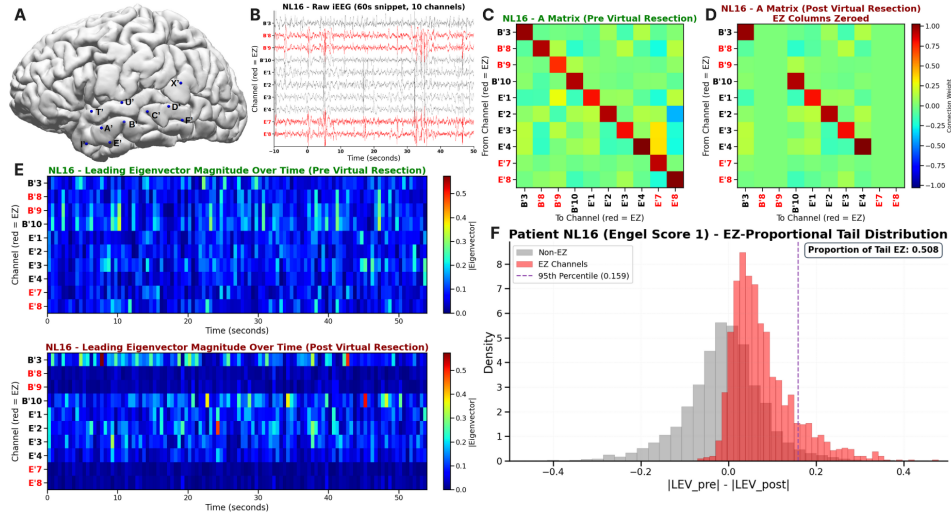


Fig. 1 Virtual Resection Pipeline for LEV-Based Surgical Outcome Prediction. Schematic overview of the computational pipeline for predicting surgical outcomes using leading eigenvector (LEV) analysis and virtual resection. **(A) Intracranial EEG Acquisition.** Stereoelectroencephalography (sEEG) electrodes are implanted according to clinical presurgical hypotheses. Raw intracranial EEG signals are recorded during interictal periods (at least 1 hour from any seizure event) at sampling rates of 500 Hz to 2 kHz. The clinical team annotates channels corresponding to the epileptogenic zone (EZ) based on ictal onset patterns and seizure semiology. EZ channels are color-coded in red. **(B) Signal Preprocessing and Segmentation.** Raw iEEG signals undergo bandpass filtering (0.5–300 Hz), notch filtering at 60 Hz and harmonics, and common average referencing. Preprocessed signals are segmented into non-overlapping 500 ms windows for model estimation. Channels not recording from gray matter or exhibiting excessive noise are excluded. **(C) Dynamic Network Model Estimation.** For each 500 ms window, a linear time-invariant state-transition matrix A is estimated using least squares regression, where $x(t+1) = Ax(t)$ describes how each channel’s activity at time t influences all channels at time $t+1$. The matrix element A_{ij} quantifies the directed influence of channel j on channel i . Only matrices with $R^2 > 0.6$ are retained. **(D) Virtual Resection.** The columns of A corresponding to EZ channels are set to zero, eliminating their outgoing influence on the network. **(E) Leading Eigenvector Computation.** The leading eigenvector (LEV) of each A matrix (for both resected and non-resected) is computed, representing the steady-state distribution of network activity. Channels with high LEV magnitude correspond to network “sinks” that accumulate influence from other regions. The change in LEV magnitude ($\Delta\text{LEV} = |\text{LEV}_{pre}| - |\text{LEV}_{post}|$) quantifies the effect of simulated EZ removal on each channel’s network influence. **(F) Feature Extraction and Outcome Prediction.** All ΔLEV values are pooled across time windows and channels. The EZ tail proportion is computed as the fraction of elements in the top 5% of the ΔLEV distribution that originate from EZ channels. This scalar feature is used in logistic regression to predict surgical outcome (Engel Class 1 vs. Engel Class 2–4). Each distribution (EZ and Non-EZ) is independently normalized so that its total area equals 1, allowing visual comparison of distribution shapes regardless of sample sizes.

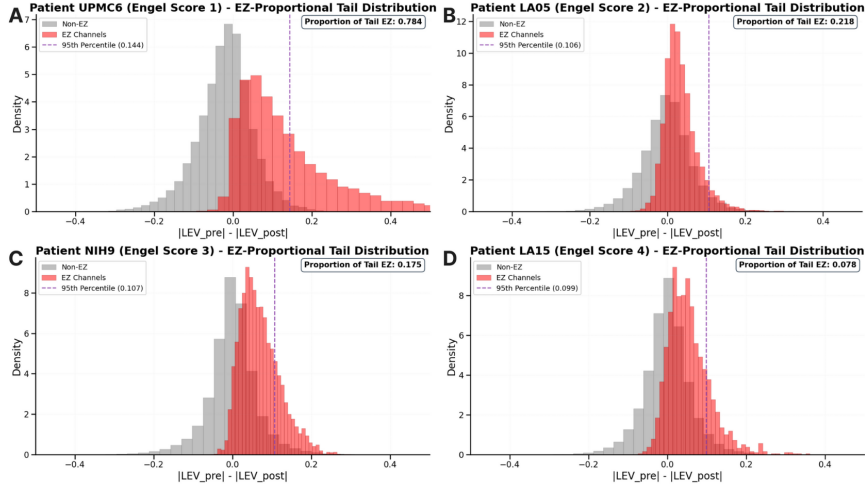


Fig. 2 Distribution of ΔLEV Values Across Patients Grouped by Engel Score. Representative histograms showing the distribution of $\Delta\text{LEV} = |\text{LEV}_{pre}| - |\text{LEV}_{post}|$ computed across all time windows and channels for individual patients. Positive values indicate reduction in eigenvector magnitude following virtual resection. The purple dashed line marks the patient-specific upper 5% quantile threshold. Red indicates elements from EZ-labeled channels and grey depicts non-EZ elements. A kernel density estimation curve (gray) overlays each histogram. The y-axis shows probability density, where each distribution (EZ and Non-EZ) is independently normalized so that its total area equals 1. This allows visual comparison of the distribution shapes regardless of the different sample sizes between EZ and Non-EZ channels. The annotation in each panel reports the proportion of tail mass attributable to EZ elements. Panels are grouped by Engel score (1–4), with patient identifiers colored by outcome (**A**) Engel 1 = seizure free; (**B–D**) Engel 2–4 = ongoing seizures. Patients achieving seizure freedom (Engel Class 1) show substantially higher EZ tail proportions compared to those with poor outcomes. Patients with worse outcomes (higher Engel scores) have progressively less of their 5% quantile threshold consisting of EZ channels.

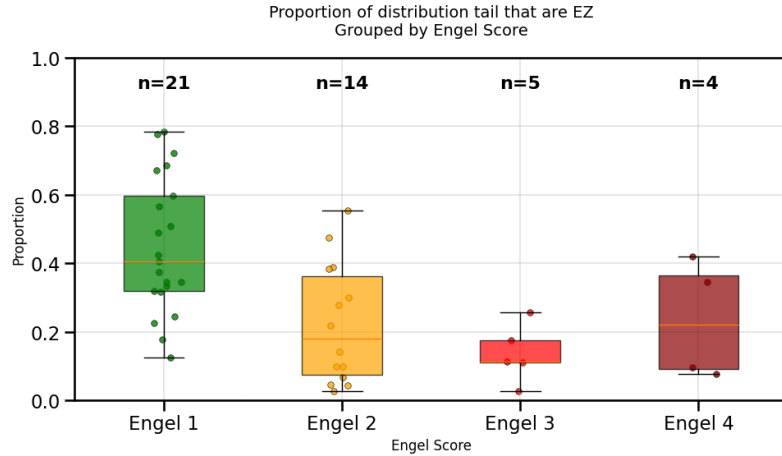


Fig. 3 EZ Tail Proportion Stratified by Engel Outcome Score. Box plots showing the distribution of EZ tail proportion values within each Engel outcome class. The EZ tail proportion is defined as the fraction of elements in the top 5% of the ΔLEV distribution that originate from clinically annotated EZ channels. Individual patients are shown as scattered points, with sample sizes annotated above each box. Box plots display median, interquartile range, and whiskers extending to $1.5 \times \text{IQR}$. Higher proportions indicate stronger effects of virtual resection on EZ channel network influence. Engel Class 1 patients (seizure free) exhibit significantly higher EZ tail proportions compared to Engel Class 2–4 patients (ongoing seizures), supporting the use of this metric as a predictive biomarker.

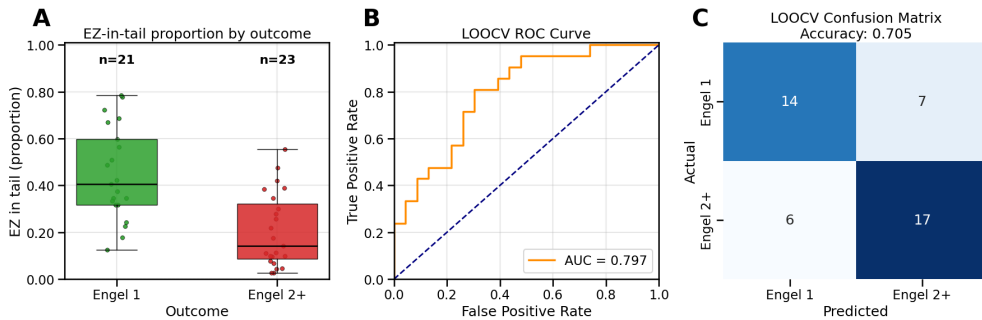


Fig. 4 Machine Learning Model Performance for Surgical Outcome Prediction. Evaluation of the logistic regression classifier using EZ tail proportion to predict surgical outcomes. **(A)** Box plots showing EZ tail proportion values in the training cohort ($n = 44$) stratified by outcome, with individual patients shown as jittered points and sample sizes annotated. Higher values are associated with Engel Class 1 outcomes. **(B)** Receiver operating characteristic (ROC) curve from leave-one-out cross-validation (LOOCV; $n = 44$), demonstrating discrimination between Engel Class 1 and Engel Class 2–4 outcomes ($\text{AUC} = 0.8$). Each prediction represents an out-of-sample probability estimate where the model was trained on all other patients. The AUC 95% confidence interval derived from 1,000 bootstrap resamples was $0.802 \pm 1.96(0.066)$ or $[0.672, 0.931]$, where 0.802 was the mean of the 1,000 bootstrap resamples and 0.066 was the standard error. The diagonal dashed line indicates chance performance. **(C)** Confusion matrix for the test set showing classification performance: overall accuracy = 71%, specificity = 74%, sensitivity = 67%. True labels are shown on the y-axis; predicted labels on the x-axis.

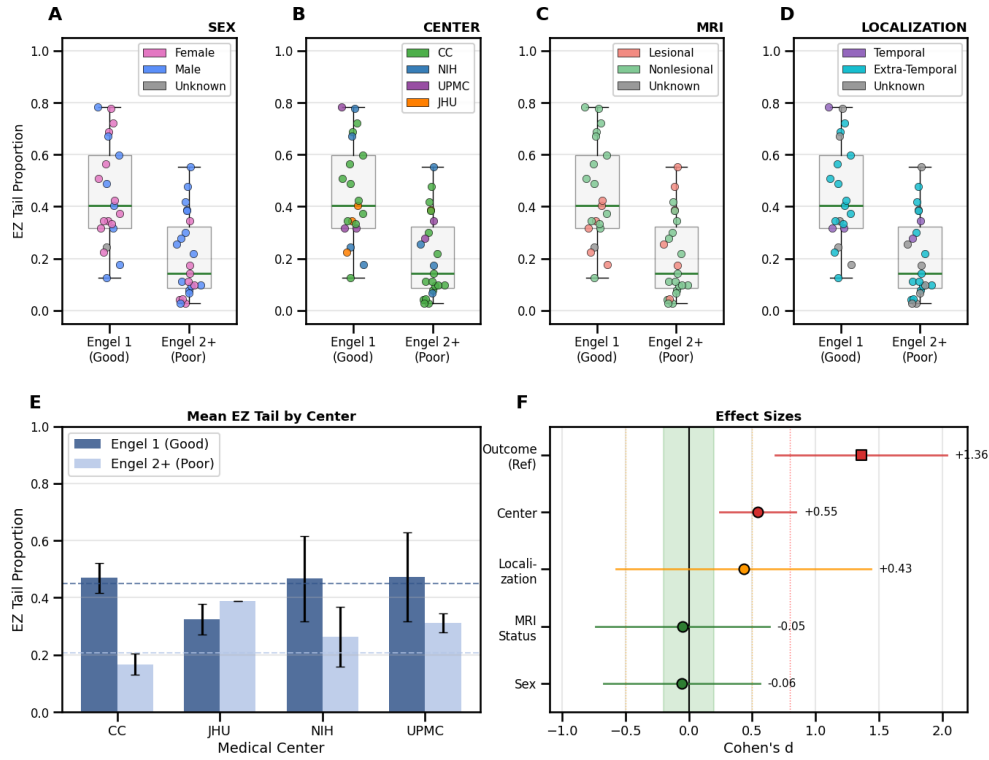


Fig. 5 Robustness of LEV-Based EZ Identification Across Clinical Subgroups. Distribution of EZ tail proportion stratified by surgical outcome (Engel I vs. Engel II+) and colored by clinical covariates, demonstrating the robustness of leading eigenvector (LEV) localization across heterogeneous patient populations. Green horizontal lines indicate group medians. **(A) Sex.** EZ tail proportion distributions show no systematic differences between male and female patients within each outcome group, indicating the LEV metric is not confounded by sex-related differences in brain network organization. **(B) Medical Center.** Distributions are consistent across four independent clinical sites (CC, NIH, UPMC, JHU), demonstrating that LEV-based EZ identification generalizes across different data acquisition systems, electrode configurations, and clinical protocols. **(C) MRI Lesional Status.** Patients with MRI-positive (lesional) and MRI-negative (nonlesional) epilepsy show comparable EZ tail proportion distributions within outcome groups, indicating that LEV localization is effective regardless of whether a structural abnormality is visible on imaging. **(D) Epilepsy Localization.** The metric performs consistently for both temporal lobe epilepsy and extra-temporal epilepsy cases, suggesting broad applicability across different epileptogenic network topologies. **(E) Mean EZ Tail by Center.** Bar chart comparing mean EZ tail proportion (\pm SEM) between Engel I (good outcome, dark blue) and Engel II+ (poor outcome, light blue) patients across medical centers. Dashed horizontal lines indicate pooled means for each outcome group. The consistent separation between outcome groups across centers supports the generalizability of the LEV metric. **(F) Effect Sizes.** Forest plot displaying Cohen's d effect sizes with 95% confidence intervals for clinical covariates. The green shaded region ($|d| < 0.2$) indicates negligible effect sizes. Sex, MRI status, and localization show negligible effects ($|d| < 0.2$), while center shows a small effect ($d = 0.55$). The reference outcome comparison (Engel I vs. Engel2+) shows a large effect ($d = 1.36$), confirming that surgical outcome—not demographic or clinical confounders—drives the observed differences in EZ tail proportion.